

1.  $\Pr(\text{disease}) = 0.01$        $\Pr(\overline{\text{disease}}) = 0.99$   
 $\Pr(\text{positive} | \text{disease}) = 0.99$        $\Pr(\overline{\text{positive}} | \text{disease}) = 0.01$   
 $\Pr(\text{positive} | \overline{\text{disease}}) = 0.03$        $\Pr(\overline{\text{positive}} | \overline{\text{disease}}) = 0.97$   
 $\Pr(\text{positive}) = \Pr(\text{positive} | \text{disease})\Pr(\text{disease}) + \Pr(\text{positive} | \overline{\text{disease}})\Pr(\overline{\text{disease}})$   
 $= (0.99)(0.01) + (0.03)(0.99) = 0.0396$   
 $\Pr(\text{disease} | \text{positive}) = \frac{\Pr(\text{positive} | \text{disease})\Pr(\text{disease})}{\Pr(\text{positive})} = \frac{(0.99)(0.01)}{0.0396} = 0.25$

2. Frequency table of the scores:

Class	Midpoint ( $m_i$ )	Frequency	Cumulative Frequency
0 – 300	150	15000	15000
300 – 500	400	50000	65000
500 – 600	550	35000	100000
600 – 700	650	45000	145000
700 – 850	775	45000	190000
850 – 1000	925	7500	197500

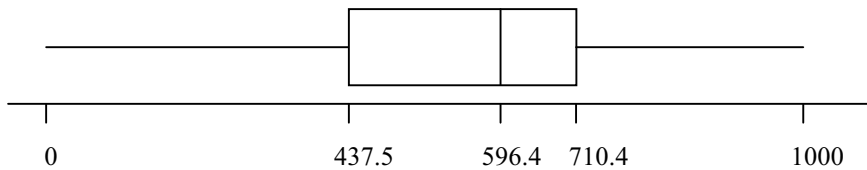
- (a)  $\text{Min} = 0$        $\text{Max} = 1000$

$$(197500)(0.25) = 49375, \quad Q_L = 300 + \frac{(200)(49375 - 15000)}{50000} = 437.5$$

$$(197500)(0.5) = 98750, \quad \text{Median} = 500 + \frac{(100)(98750 - 65000)}{35000} = 596.4$$

$$(197500)(0.75) = 148125, \quad Q_U = 700 + \frac{(150)(148125 - 145000)}{45000} = 710.4$$

Boxplot :



$$(b) \mu = \frac{1}{197500} \sum_{i=1}^k m_i f_i = \frac{112562500}{197500} = 569.9$$

$$(c) \Pr(500 \leq \bar{X} \leq 600) = \Pr\left(\frac{500 - 569.9}{207/\sqrt{50}} \leq \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \leq \frac{600 - 569.9}{207/\sqrt{50}}\right)$$

$$\approx \Phi(1.03) - \Phi(-2.39) = 0.8485 - (1 - 0.9916) = 0.8401$$

$$(d) (197500)(0.95) = 187625$$

$$95^{\text{th}} \text{ percentile} = 700 + \frac{(150)(187625 - 145000)}{45000} = 842.1$$

3. (a)  $\lambda = 1.5$        $t = 12$   
 $\lambda t = (1.5)(12) = 18$

(b) Let  $X$  be the number of failures in a year. Then  $X \sim \mathcal{P}(18)$ .

$$\begin{aligned}\Pr(|X - E(X)| > \sqrt{\text{Var}(X)}) &= 1 - \Pr(|X - 18| \leq \sqrt{18}) \\ &= 1 - \Pr(18 - \sqrt{18} \leq X \leq 18 + \sqrt{18}) \\ &= 1 - \Pr(13.757 \leq X \leq 22.243) \\ &= 1 - \Pr(14 \leq X \leq 22) \\ &= 1 - \sum_{i=14}^{22} \frac{e^{-18}(18)^i}{i!} \\ &= 1 - 0.7125 = 0.2875\end{aligned}$$

(c)  $p = \Pr(\text{monthly loss less than 1000})$

$$= \Pr(\text{number of monthly failures} \leq 2) = \sum_{i=0}^2 \frac{e^{-1.5}(1.5)^i}{i!} = 0.8088$$

Let  $Y$  be the number of months during the year with monthly loss less than 1000.  
Then  $Y \sim b(12, 0.8088)$ .

$$\Pr(Y \geq 8) = \sum_{i=8}^{12} \binom{12}{i} (0.8088)^i (0.1912)^{12-i} = 0.9386$$

(d)  $T \sim \Gamma(25, 1.5)$

4. (a) A 95% confidence for  $p$  is

$$\hat{p} \pm Z_{0.025} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} = \frac{13}{30} \pm (1.96) \sqrt{\frac{(13/30)(17/30)}{30}} = 0.4333 \pm 0.1773 = [25.6\%, 61.06\%]$$

(b) Although the sample proportion is more than 40%, the population proportion may not be more than 40% because the sampling error is 17.73% which is very large. Indeed the range of the confidence interval is so wide that the population proportion can be as low as 25.6% or as high as 61.6%. The newspaper had over-interpreted the experimental result in the statement.

$$(c) n = \frac{Z_{0.025}^2 p(1-p)}{D^2} = \frac{(1.96)^2 (13/30)(17/30)}{(0.025)^2} = 1509.32 \approx 1510$$

(OR use the conservative sample size :

$$n = \frac{Z_{0.025}^2}{4D^2} = \frac{(1.96)^2}{4(0.025)^2} = 1536.64 \approx 1537)$$

5. (a) Test  $H_0 : \sigma_1^2 = \sigma_2^2$  vs  $H_1 : \sigma_1^2 \neq \sigma_2^2$  at  $\alpha = 0.1$ .

$$\text{Test statistic: } F = \frac{S_1^2}{S_2^2}$$

$$\text{Reject } H_0 \text{ if } F > F(7,7,0.05) = 3.79 \text{ or } F < F(7,7,0.95) = \frac{1}{3.79} = 0.2639.$$

$$\text{From the data, } \bar{X} = 2577.75, S_1^2 = 8158.79, \bar{Y} = 2550.5, S_2^2 = 5485.43.$$

$$F_{obs} = \frac{8158.79}{5485.43} = 1.4874 \Rightarrow 0.2639 < F_{obs} < 3.79$$

Hence do not reject  $H_0$  at  $\alpha = 0.1$ . The data does not provide enough evidence that the population variances are unequal.

- (b) Test  $H_0 : \mu_1 = \mu_2$  vs  $H_1 : \mu_1 \neq \mu_2$  at  $\alpha = 0.05$ .

Assume equal population variances.

$$\text{Test statistic: } T = \frac{\bar{X} - \bar{Y}}{S_{pool} \sqrt{\frac{1}{8} + \frac{1}{8}}}$$

$$\text{Reject } H_0 \text{ if } |T| > t_{14,0.025} = 2.145.$$

$$\text{From the data, } \bar{X} = 2577.75, S_1^2 = 8158.79, \bar{Y} = 2550.5, S_2^2 = 5485.43.$$

$$S_{pool}^2 = \frac{(7)(8158.79) + (7)(5485.43)}{7+7} = 6822.11$$

$$|T_{obs}| = \frac{|2577.75 - 2550.5|}{\sqrt{(6822.11) \left( \frac{1}{8} + \frac{1}{8} \right)}} = 0.6598 < 2.145$$

Hence do not reject  $H_0$  at  $\alpha = 0.05$ . The data does not provide enough evidence that in average the measurements from the two testers are difference.

- (c) A 95% confidence interval for  $\mu_1$  is

$$\bar{X} \pm t_{7,0.025} \frac{S_1}{\sqrt{8}} = 2577.75 \pm (2.365) \sqrt{\frac{8158.79}{8}} = 2577.75 \pm 75.53 = [2502.22, 2653.28]$$

A 95% confidence interval for  $\mu_2$  is

$$\bar{Y} \pm t_{7,0.025} \frac{S_2}{\sqrt{8}} = 2550.5 \pm (2.365) \sqrt{\frac{5485.43}{8}} = 2550.5 \pm 61.93 = [2488.57, 2612.43]$$

(OR use the pooled sample variance,

$$\bar{X} \pm t_{14,0.025} \frac{S_{pool}}{\sqrt{8}} = 2577.75 \pm (2.145) \sqrt{\frac{6822.11}{8}} = 2577.75 \pm 62.64 = [2515.11, 2640.39]$$

$$\bar{Y} \pm t_{14,0.025} \frac{S_{pool}}{\sqrt{8}} = 2550.5 \pm (2.145) \sqrt{\frac{6822.11}{8}} = 2550.5 \pm 62.64 = [2487.86, 2613.14] )$$

- (d) Let A be the event that the first interval contains  $\mu_1$ , B be the event that the second interval contains  $\mu_2$ . Then the probability that both intervals will contain the corresponding parameters is

$$\Pr(A \cap B) = 1 - \Pr(\bar{A} \cup \bar{B}) \geq 1 - \Pr(\bar{A}) - \Pr(\bar{B}) = 1 - 0.05 - 0.05 = 0.9.$$

Therefore we only know that the overall confidence is at least 90%. However, we don't know whether the overall confidence can be 95% or not.

- (e) A 90% confidence interval for  $\mu_1 - \mu_2$  is given by

$$\begin{aligned} (\bar{X} - \bar{Y}) \pm t_{14,0.05} S_{pool} \sqrt{\frac{1}{8} + \frac{1}{8}} &= (2577.75 - 2550.5) \pm (1.761) \sqrt{(6822.11) \left( \frac{1}{8} + \frac{1}{8} \right)} \\ &= 27.25 \pm 72.73 = [-45.48, 99.98] \end{aligned}$$

6. (a)  $H_0 : p_V \geq p_N$  vs  $H_0 : p_V < p_N$ .

(b) Test statistic :

$$Z = \frac{\hat{p}_V - \hat{p}_N}{\sqrt{\hat{p}(1-\hat{p}) \left( \frac{1}{n_V} + \frac{1}{n_N} \right)}}$$

Reject  $H_0$  if  $Z < -Z_{0.05} = -1.645$ .

From the data,  $\hat{p}_V = \frac{9}{51} = 0.1765$ ,  $\hat{p}_N = \frac{17}{45} = 0.3778$ ,  $\hat{p} = \frac{9+17}{51+45} = 0.2708$

$$Z_{obs} = \frac{0.1765 - 0.3778}{\sqrt{(0.2708)(0.7292) \left( \frac{1}{51} + \frac{1}{45} \right)}} = -2.2149 < -1.645.$$

Reject  $H_0$  at  $\alpha = 0.05$ . The data provided strong evidence that the vaccination is effective in protecting the laboratory animals from that disease.

- (c) A 95% confidence interval for  $p_V - p_N$  is given by

$$\begin{aligned} (\hat{p}_V - \hat{p}_N) \pm Z_{0.025} \sqrt{\frac{\hat{p}_V(1-\hat{p}_V)}{n_V} + \frac{\hat{p}_N(1-\hat{p}_N)}{n_N}} \\ = (0.1765 - 0.3778) \pm (1.96) \sqrt{\frac{(0.1765)(0.8235)}{51} + \frac{(0.3778)(0.6222)}{45}} \\ = -0.2013 \pm 0.1761 = [-0.3774, -0.0252] \end{aligned}$$

Since all the possible values in this interval are less than zero, we have high confidence that  $p_V$  is less than  $p_N$ , i.e. the vaccination is effective.